

	Mean	Std Dev	Z-Score	Assumption of CI	100(1- α)% CI
Sampling distribution of \bar{x}	$\mu_{\bar{x}} = \mu$	$\sigma_{\bar{x}} = \frac{\sigma}{\sqrt{n}}$	$z = \frac{\bar{x} - \mu_{\bar{x}}}{\sigma_{\bar{x}}} = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}}$ $t = \frac{\bar{X} - \mu}{S/\sqrt{n}}$	Large Sample: Random; n>30	for μ : $\bar{x} \pm z_{\alpha/2} \sigma_{\bar{x}} = \bar{x} \pm z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$ if σ unknown, n>30, $\bar{x} \pm z_{\alpha/2} \frac{s}{\sqrt{n}}$
				Small Sample: Random; apx. Normal	If σ unknown, n<30: $\bar{x} \pm t_{\alpha/2} \frac{s}{\sqrt{n}}$, df=n-1
Sampling distribution of \hat{p}	$\hat{p} = p$	$\sigma_{\hat{p}} = \sqrt{pq/n}$ 0 or 1 $\notin (\hat{p} - 3\hat{\sigma}_{\hat{p}}, \hat{p} + 3\hat{\sigma}_{\hat{p}})$ $\hat{p} \sim N\left(\mu_{\hat{p}} = p, \sigma_{\hat{p}}^2 = \frac{pq}{n}\right)$		Random 0 or 1 $\notin (\hat{p} - 3\hat{\sigma}_{\hat{p}}, \hat{p} + 3\hat{\sigma}_{\hat{p}})$	$\hat{p} \pm z_{\alpha/2} \sigma_{\hat{p}} = \hat{p} \pm z_{\alpha/2} \sqrt{pq/n} \approx \hat{p} \pm z_{\alpha/2} \sqrt{\hat{p}\hat{q}/n}$
				When n is extremely large	$\tilde{p} \pm z_{\alpha/2} \sqrt{\frac{\tilde{p}(1-\tilde{p})}{n+4}}$ where $\tilde{p} = \frac{x+2}{n+4}$

Finding sample size for estimating population mean:
$$n = \frac{(Z_{\alpha/2})^2 \sigma^2}{SE^2}$$

Finding Sample size for estimating population proportion
$$n = \frac{(Z_{\alpha/2})^2 (pq)}{SE^2}$$

Converting a CI to accommodate a finite population:

$$\bar{x} \pm z_{\alpha/2} \cdot \left[\frac{\sigma}{\sqrt{n}} \cdot \sqrt{\frac{N-n}{N}} \right], \quad \bar{x} \pm t_{\alpha/2} \cdot \left[\frac{s}{\sqrt{n}} \cdot \sqrt{\frac{N-n}{N}} \right], \quad \hat{p} \pm z_{\alpha/2} \cdot \left[\sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \cdot \sqrt{\frac{N-n}{N}} \right] \quad \text{use when } n/N > 0.05$$

Commonly Used Value of $z_{\alpha/2}$

Confidence Level 100(1- α)	α	$\alpha/2$	$z_{\alpha/2}$
90%	0.10	0.05	1.645
95%	0.05	0.025	1.96
99%	0.01	0.005	2.575

Reduce confidence level, narrower CI.

Rejection Regions for Common Value of α

α	Lower Tailed	Upper Tailed	Two Tailed
0.10	$z < -1.28$	$z > 1.28$	$z < -1.645$, or $z > 1.645$
0.05	$z < -1.645$	$z > 1.645$	$z < -1.96$, or $z > 1.96$
0.01	$z < -2.33$	$z > 2.33$	$z < -2.575$, or $z > 2.575$

	Assumption	Test Statistic	Alternative	Reject H_0 if	p- value
Large Sample: Testing a Population Mean	Normal, or if n is large (≥ 30),	if σ unknown while n is large, replace $z = \frac{\bar{x} - \mu}{s/\sqrt{n}}$	$H_a : \mu > \mu_0$	$z > z_\alpha$	Area under N(0,1) curve right of z
			$H_a : \mu < \mu_0$	$z < -z_\alpha$	Area under N(0,1) curve left of z
			$H_a : \mu \neq \mu_0$	$ z > z_{\alpha/2}$	Twice area under N(0,1) curve right of $ z $, $z > z_{\alpha/2}$ or $z < -z_{\alpha/2}$
Small Sample: Testing a Population Mean	Random Apx. Normal	$\frac{\bar{X} - \mu}{S / \sqrt{n}} \sim t_{n-1}$	$H_a : \mu > \mu_0$	$t > t_\alpha$	Area under t curve right of t
			$H_a : \mu < \mu_0$	$t < -t_\alpha$	Area under t curve left of t
			$H_a : \mu \neq \mu_0$	$ t > t_{\alpha/2}$	Twice area under t curve right of $ t $ $t > t_{\alpha/2}$ or $t < -t_{\alpha/2}$
Large-Sample: Testing a Population Proportion	Random $p_0 \pm 3\sigma_{\hat{p}} \approx p_0 \pm 3\sqrt{\frac{p_0(1-p_0)}{n}} \in (0,1)$	$z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}} \sim N(0,1)$	$H_a : p > p_0$	$z > z_\alpha$	Area under N(0,1) curve right of z
			$H_a : p < p_0$	$z < -z_\alpha$	Area under N(0,1) curve left of z
			$H_a : p \neq p_0$	$ z > z_{\alpha/2}$	Twice area under N(0,1) curve right of $ z $ $z > z_{\alpha/2}$ or $z < -z_{\alpha/2}$
Testing a Population Variance	Random Apx. Normal	$\chi^2 = \frac{(n-1)s^2}{\sigma_0^2} \sim \chi_{n-1}^2$	$H_a : \sigma^2 > \sigma_0^2$	$\chi^2 > \chi_\alpha^2$	100(1-α)% confidence interval for σ^2: $\left(\frac{(n-1)s^2}{\chi_{\alpha/2}^2}, \frac{(n-1)s^2}{\chi_{1-\alpha/2}^2} \right)$
			$H_a : \sigma^2 < \sigma_0^2$	$\chi^2 < \chi_{1-\alpha}^2$	
			$H_a : \sigma^2 \neq \sigma_0^2$	$\chi^2 > \chi_{\alpha/2}^2$ or $\chi^2 < \chi_{1-\alpha/2}^2$	

	Assumption	Test Statistic	CI	Alternative	Reject H_0 if	p- value
Large Sample: Testing a Population Mean Difference	Randomly selected in independent manner; Sample size sufficient large	$z = \frac{\bar{x}_1 - \bar{x}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}} \sim N(0,1)$	$(\bar{x}_1 - \bar{x}_2) \pm z_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$ If Std unknown while n>30 $(\bar{x}_1 - \bar{x}_2) \pm z_{\alpha/2} \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$	$\mu_1 - \mu_2 > D_0$	$z > z_{\alpha}$	Area under N(0,1) curve right of z
		if σ unknown while n is large, replace		$\mu_1 - \mu_2 < D_0$	$z < -z_{\alpha}$	Area under N(0,1) curve left of z
		$z = \frac{\bar{x}_1 - \bar{x}_2 - (\mu_1 - \mu_2)}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} \sim N(0,1)$		$\mu_1 - \mu_2 \neq D_0$	$ z > z_{\alpha/2}$	Twice area under N(0,1) curve right of z , $z > z_{\alpha/2}$ or $z < -z_{\alpha/2}$
Small Sample: Testing a Population Mean Difference	Random, Independent, Apx. Normal $\sigma_1 = \sigma_2$	$t = \frac{\bar{x}_1 - \bar{x}_2 - (\mu_1 - \mu_2)}{s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim t_{n_1+n_2-2}$	$\bar{x}_1 - \bar{x}_2 \pm t_{\alpha/2} s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$ Where $\sim t_{n_1+n_2-2}$	$\mu_1 - \mu_2 > D_0$	$t > t_{\alpha}$	Area under t curve right of t
		$s_p^2 = \frac{(n_1-1)s_1^2 + (n_2-1)s_2^2}{n_1+n_2-2}$		$\mu_1 - \mu_2 < D_0$	$t < -t_{\alpha}$	Area under t curve left of t
				$\mu_1 - \mu_2 \neq D_0$	$ t > t_{\alpha/2}$	Twice area under t curve right of t $t > t_{\alpha/2}$ or $t < -t_{\alpha/2}$
	Random, independent $N(\mu_i, \sigma_i^2), i = 1, 2$ & $\sigma_1^2 \neq \sigma_2^2$	$t = \frac{(\bar{x}_1 - \bar{x}_2) - D_0}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} \sim t_{df}$ where $df = \frac{(s_1^2/n_1 + s_2^2/n_2)^2}{\frac{(s_1^2/n_1)^2}{n_1-1} + \frac{(s_2^2/n_2)^2}{n_2-1}}$				
Paired-Observation Comparisons: Comparing Two Means When the Samples are Dependent	Random sample from Single normal population of differences	$z = \frac{\bar{d} - \mu_d}{\sigma_d/\sqrt{n}} \sim N(0,1)$	$\bar{d} \pm z_{\alpha/2} \frac{\sigma_d}{\sqrt{n}}$ $\bar{d} \pm t_{\alpha/2} \frac{s_d}{\sqrt{n}}$	$\mu_d > D_0$	$t > t_{\alpha}$	Area under t curve right of t
		$t = \frac{\bar{d} - \mu_d}{s_d/\sqrt{n}} \sim t_{n-1}$		$\mu_d < D_0$	$t < -t_{\alpha}$	Area under t curve left of t
				$\mu_d \neq D_0$	$ t > t_{\alpha/2}$	Twice area under t curve right of t $t > t_{\alpha/2}$ or $t < -t_{\alpha/2}$

	Test Statistic	CI	Alternative	Reject H_0 if	p- value	
Large-Sample: Testing a Population Proportion Difference	$z = \frac{(\hat{p}_1 - \hat{p}_2) - (p_1 - p_2)}{\sqrt{\left(\frac{p_1 q_1}{n_1} + \frac{p_2 q_2}{n_2}\right)}} \sim N(0,1)$	$(\hat{p}_1 - \hat{p}_2) \pm z_{\alpha/2} \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}$	Adjusted CI for very small n: $\tilde{p}_1 - \tilde{p}_2 \pm z_{\alpha/2} \sqrt{\frac{\tilde{p}_1(1-\tilde{p}_1)}{n_1+2} + \frac{\tilde{p}_2(1-\tilde{p}_2)}{n_2+2}}$ where $\tilde{p}_i = \frac{x_i + 1}{n_i + 2}$, for $i=1,2$.			
	$z = \frac{(\hat{p}_1 - \hat{p}_2) - (p_1 - p_2)}{\sqrt{\left(\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}\right)}} \sim N(0,1)$ <p style="text-align: center;">Assumption: Two Samples are randomly selected in an independent manner; Samples sizes are both large so that the sampling distribution of the proportion difference will be approximately normal</p>		$H_0: p_1 - p_2 = 0 \text{ (i.e. } D_0 = 0)$ $z = \frac{(\hat{p}_1 - \hat{p}_2) - 0}{\sqrt{\hat{p}(1-\hat{p})\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$ <p>Test Statistic:</p>	$p_1 - p_2 > D_0$	$z > z_{\alpha}$	Area under N(0,1) curve right of z
				$p_1 - p_2 < D_0$	$z < -z_{\alpha}$	Area under N(0,1) curve left of z
		$H_0: p_1 - p_2 = D_0 \text{ (i.e. } D_0 \neq 0)$ $z = \frac{(\hat{p}_1 - \hat{p}_2) - D_0}{\sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}}}$ <p style="text-align: center;">OR</p> $\text{where } \hat{p} = \frac{n_1 \hat{p}_1 + n_2 \hat{p}_2}{n_1 + n_2}$	$p_1 - p_2 \neq D_0$	$ z > z_{\alpha/2}$	Twice area under N(0,1) curve right of z $z > z_{\alpha/2}$ or $z < -z_{\alpha/2}$	
Valid F-Test for Equal Variance	<p style="text-align: center;">Assumption: Both Sampled Populations are normally distributed; The samples are random and independent.</p>	$F = \frac{s_1^2}{s_2^2} \sim F \text{ with } n_1 - 1 \text{ and } n_2 - 1 \text{ dfs}$	$H_a: \sigma_1^2 > \sigma_2^2$	$F > F_{\alpha}$ with $n_1 - 1$ and $n_2 - 1$ dfs p-value $< \alpha$		
		$F = \frac{s_2^2}{s_1^2} \sim F \text{ with } n_2 - 1 \text{ and } n_1 - 1 \text{ dfs}$	$H_a: \sigma_1^2 < \sigma_2^2$	$F > F_{\alpha}$ with $n_2 - 1$ and $n_1 - 1$ dfs p-value $< \alpha$		
		$F = \frac{\text{larger of } s_1^2 \text{ and } s_2^2}{\text{smaller of } s_1^2 \text{ and } s_2^2} \sim F$ <p style="text-align: center;">where</p> $df_1 = \{\text{size of sample with larger variance}\} - 1$ $df_2 = \{\text{size of sample with smaller variance}\} - 1$	$H_a: \sigma_1^2 \neq \sigma_2^2$	$F > F_{\alpha/2}$ with df_1 and df_2 dfs p-value $< \alpha$		

Finding sample size for estimating population mean/proportion difference

$$n = \frac{(z_{\alpha/2})^2 (\sigma_1^2 + \sigma_2^2)}{e^2}, \text{ and } n = \frac{(z_{\alpha/2})^2 (p_1 q_1 + p_2 q_2)}{e^2}$$